

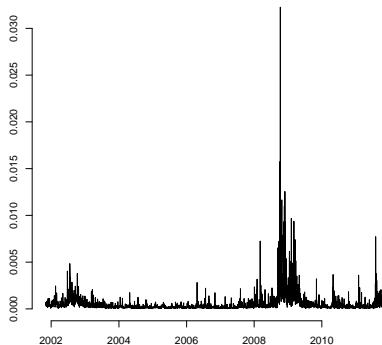
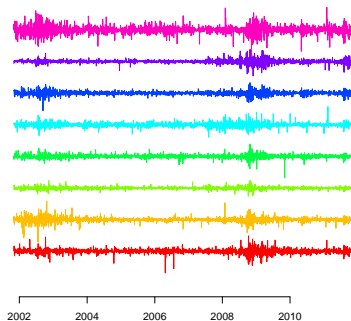
Hierarchical Gaussian Graphical Models: Beyond Reversible Jump

Alex Lenkoski

Institut für Angewandte Mathematik
Universität Heidelberg

A Troubling Dataset

Daily returns on 20 assets from 2002 to 2012



The G-Wishart Distribution

Suppose $\mathcal{D} = \{\mathbf{Z}^{(1)}, \dots, \mathbf{Z}^{(n)}\}$ with

$$\mathbf{Z}^{(j)} \sim \mathcal{N}_p(0, \mathbf{K}^{-1}).$$

We consider the G-Wishart prior $\mathcal{W}_G(\delta, \mathbf{D})$

$$pr(\mathbf{K} | G, \delta, \mathbf{D}) = \frac{1}{l_G(\delta, \mathbf{D})} |\mathbf{K}|^{(\delta-2)/2} \exp\left(-\frac{1}{2} \langle \mathbf{K}, \mathbf{D} \rangle\right) \mathbf{1}_{\mathbf{K} \in P_G},$$

where

$$\mathbf{K} \in P_G \Rightarrow K_{ij} = 0, \quad (i, j) \notin G.$$

Then

$$\mathbf{K} | \mathcal{D} \sim \mathcal{W}_G(\delta + n, \mathbf{D}^*)$$

where $\mathbf{D}^* = \mathbf{D} + \sum \mathbf{Z}^{(j)} \mathbf{Z}^{(j)'}$.

Cholesky Basics

Let Φ be the Cholesky decomposition of $\mathbf{K} \sim \mathcal{W}_G(\delta, \mathbf{D})$. Then

$$pr(\Phi|G) = \prod_{i=1}^p \Phi_{ii}^{(\delta + \nu_i^G - 1)} \exp\left(-\frac{1}{2} \langle \Phi' \Phi, \mathbf{D} \rangle\right)$$

Let F be the fill-in graph associated with G , thus

$$\Phi_{ij} = 0 \quad (i, j) \notin F$$

and

$$\Phi_{ij} = -\frac{1}{\Phi_{ii}} \sum_{l=1}^i \Phi_{li} \Phi_{lj} \quad (i, j) \in F \setminus G.$$

Finally, if $\mathbf{K} \sim \mathcal{W}_G(\delta, \mathbb{I}_p)$, then

$$pr(\Phi|G) = \prod_{i=1}^p \Phi_{ii}^{\delta + \nu_i^G - 1} \exp\left(-\frac{1}{2} \sum_{(i,j) \in F} \Phi_{ij}^2\right).$$

Reversible Jump

Key result of Dobra, Lenkoski and Rodriguez (2011):

$$R_g^+ = \frac{pr(\mathcal{D}|\mathbf{K}')}{pr(\mathcal{D}|\mathbf{K}^{[s]})} \frac{pr(\mathbf{K}'|G')}{pr(\mathbf{K}^{[s]}|G^{[s]})} \times \\ \times \frac{J(\mathbf{K}' \rightarrow (\Phi')^{\nu(G')})}{J(\mathbf{K}^{[s]} \rightarrow (\Phi^{[s]})^{\nu(G^{[s]})})} \frac{J\left(\left((\Phi^{[s]})^{\nu(G^{[s]})}, \gamma\right) \rightarrow (\Phi')^{\nu(G')}\right)}{\frac{1}{\sigma_g \sqrt{2\pi}} \exp\left(-\frac{(\Phi'_{i_0 j_0} - \Phi^{[s]}_{i_0 j_0})^2}{2\sigma_g^2}\right)}.$$

Famous Last Words:

"Basically, at this point, I feel like this problem has been solved"

– A.L., Zürich, 2011

Conditional Bayes Factors (CBFs)

Let $f = (p - 1, p)$ and $\Phi^{-f} = \Phi \setminus \{\Phi_{p-1,p}, \Phi_{p,p}\}$. Suppose that $G \subset G'$ with $G' \setminus G = f$.

Instead of considering

$$\frac{\text{pr}(\mathcal{D}|G')}{\text{pr}(\mathcal{D}|G)}$$

we consider the CBF

$$\frac{\text{pr}(\mathcal{D}, \Phi^{-f}|G')}{\text{pr}(\mathcal{D}, \Phi^{-f}|G)}$$

and note

$$\begin{aligned}\text{pr}(\mathcal{D}, \Phi^{-f}|G) &= \int_0^{+\infty} \text{pr}(\mathcal{D}|\Phi)\text{pr}(\Phi|G)d\Phi_{pp} \\ \text{pr}(\mathcal{D}, \Phi^{-f}|G') &= \int_0^{+\infty} \int_{-\infty}^{+\infty} \text{pr}(\mathcal{D}|\Phi)\text{pr}(\Phi|G')d\Phi_f d\Phi_{pp}.\end{aligned}$$

Really? (A Humbling Calculation)

Let $a = p - 1$, we have that

$$pr(\mathcal{D}, \Phi^{-f} | G) \propto \frac{1}{I_G} \exp\left(-\frac{1}{2}(2D_{ap}^* \phi_{aa} \phi_0 + D_{pp}^* \phi_0^2)\right)$$

where ϕ_0 is the fill-in from Φ^{-f} of f . Likewise,

$$pr(\mathcal{D}, \Phi^{-f} | G') \propto \frac{\Phi_{aa}}{I_{G'}} \int_{-\infty}^{+\infty} \exp\left(-\frac{1}{2}(2D_f^* \Phi_{aa} \Phi_f + D_{pp}^* \Phi_f^2)\right) d\Phi_f.$$

Thus,

$$\frac{pr(\mathcal{D}, \Phi^{-f} | G')}{pr(\mathcal{D}, \Phi^{-f} | G)} = \Phi_{aa} \left(\frac{2\pi}{D_{pp}^*}\right)^{1/2} \exp\left(\frac{1}{2} D_{pp}^* (\phi_0 - \mu)^2\right) \frac{I_G}{I_{G'}}$$

where $\mu = -\Phi_{aa} D_f^* / D_{pp}^*$.

Avoiding Normalizing Constants

We thus have

$$\frac{pr(G|\mathcal{D}, \Phi^{-f})}{pr(G'|\mathcal{D}, \Phi^{-f})} = H(\Phi^{-f}, \mathbf{D}^*) \frac{I_{G'}}{I_G} \frac{pr(G')}{pr(G)}.$$

Suppose that $\tilde{\mathbf{K}} \sim \mathcal{W}_{G'}(\delta, \mathbf{D})$. The exchange algorithm (Murray et al. 2006) allows us to replace this with

$$\frac{H(\Phi^{-f}, \mathbf{D}^*) pr(G')}{H(\tilde{\Phi}^{-f}, \mathbf{D}) pr(G)}$$

The double Metropolis Hastings (Liang 2010) algorithm allows $\tilde{\mathbf{K}}$ to be sampled via an MCMC step from \mathbf{K} .

A perfectly terrible update

Under the prior $\mathcal{W}_G(\delta, \mathbb{I}_p)$ and current state \mathbf{K} , consider the following algorithm

1. Determine Φ from \mathbf{K}
2. Propose Ψ :
 - a. Sample $c \sim \chi_{\delta + \nu_j^G}^2$ and set $\Psi_{ii} = c^{1/2}$
 - b. Sample $\Psi_{ij} \sim \mathcal{N}(0, 1)$ for $(i, j) \in G$
 - c. Complete Ψ for all $(i, j) \in F \setminus G, i < j$
3. Compute

$$\alpha = \exp \left(-\frac{1}{2} \sum_{(i,j) \in F \setminus G} (\Psi_{ij}^2 - \Phi_{ij}^2) \right)$$

4. With probability $\min\{\alpha, 1\}$ set $\mathbf{K} = \Psi' \Psi$

A terrible posterior sampler, but perfect here!

Simulation Results

	Time (sec)		MSE	
	Mean	SD	Mean	SD
CL	182.5	(4.1)	0.0088	(6e-04)
RJ	1518.4	(30.2)	0.0349	(0.0025)

All of this, and in six dimensions!

A Multivariate Graphical Stochastic Volatility Model

We can now address our “troubling” dataset.

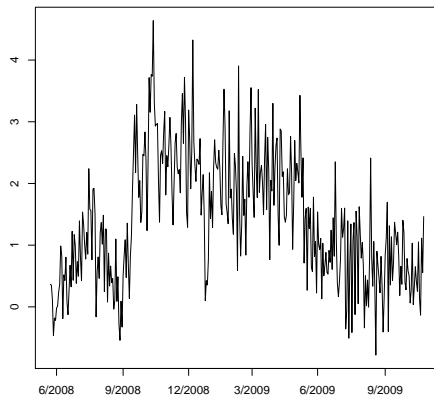
$$\begin{aligned}\mathbf{Y}_t | \mathbf{K}, X_t &\sim \mathcal{N}_p(\mathbf{0}, \exp(X_t) \mathbf{K}^{-1}) \\ X_t | \phi, X_{t-1}, \beta, \tau &\sim \mathcal{N}(\phi X_{t-1} + \beta \text{VIX}_{t-1}, \tau^{-1}) \\ \mathbf{K} | G &\sim \mathcal{W}_G(\delta, \mathbb{I}_p) \\ G &\sim pr(G)\end{aligned}$$

And note that

$$\mathbf{K} | \{\mathbf{Y}^{(1)}, \dots, \mathbf{Y}^{(T)}\}, \mathbf{X}, G \sim \mathcal{W}_G \left(\delta + T, \mathbf{D} + \sum_{t=1}^T \frac{\mathbf{Y}^{(t)} \mathbf{Y}^{(t)'}}{\exp(X_t)} \right)$$

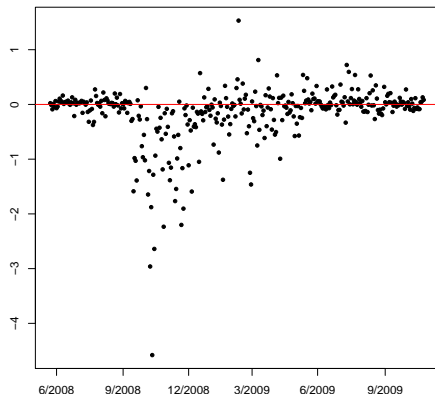
Volatility Components

Posterior mean of $X_{t+1} | \mathbf{Y}_1, \dots, \mathbf{Y}_t$



Predictive Performance

We now forecast one step ahead $pr(\mathbf{Y}^{(t+1)}|\mathbf{Y}^{(1:t)})$, comparing the volatility model to one without heteroskedasticity



Above, the difference in Energy Scores.

Conclusions

Paper on the arXiv

- ▶ Yuan Cheng and Alex Lenkoski. *A Multivariate Graphical Stochastic Volatility Model*, in submission

More work on CBFs

- ▶ Anna Karl and Alex Lenkoski. *Instrumental Variable Bayesian Model Averaging via Conditional Bayes Factors*, in submission
- ▶ Alexander Jordan and Alex Lenkoski. *Tobit Bayesian Model Averaging and Foreign Direct Investment*, in submission
- ▶ Tobias Kaiser and Alex Lenkoski. *Bayesian Verification of Instrument Validity via Conditional Bayes Factors*, in submission